

Construction and validation of a scale of stigma against individuals diagnosed with schizophrenia

Construcción y validación de una escala de estigma hacia personas diagnosticadas con esquizofrenia

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Abstract:

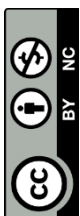
Antecedents: Individuals diagnosed with schizophrenia are not just dealing with their diagnoses. They are facing stigma due to their pathology. International research has proposed that individuals diagnosed with schizophrenia suffer more stigma than other types of mental issues. However, in Chile, a valid scale is not available to measure stigma against individuals diagnosed with schizophrenia. **Objectives:** To fill this gap, this research is aimed to develop and validate a scale to measure stigma against individuals diagnosed with schizophrenia. **Methods:** Two stages were completed to achieve the research objective. First, a pool of items was developed based on the three critical components of stigma, cognitive, affective and behavioural. Three independent judges were asked to assess the content aspects of the content validity of the items. Second, following an instrumental and longitudinal design with non-probabilistic with a quota sampling by gender (N = 607) the validity and reliability of the final scale was assessed. **Results:** A one-dimensional scale composed of 22 items showed good statistical boundaries. The observed factor loadings suggest that the items adequately represent the dimension ($\lambda > .6$), and the reliability estimates are optimal ($\alpha > .8$; $\omega > .8$). Results suggest that the scale can be used the respondents' gender irrespectively.

Keywords: stigma; schizophrenia; scale validation.

Resumen:

Antecedentes: Las personas diagnosticadas con esquizofrenia no solo enfrentan su diagnóstico. Ellos también tienen que enfrentar el estigma producto de su patología. Investigaciones internacionales han propuesto que las personas diagnosticadas con esquizofrenia sufren más de estigma que otras patologías mentales. Sin embargo, en Chile no existe una escala validada para medir el estigma hacia personas diagnosticadas con esquizofrenia. **Objetivos:** Para cubrir esta necesidad en la literatura, esta investigación tiene como objetivo desarrollar y validar un instrumento para medir el estigma hacia personas diagnosticadas con esquizofrenia. **Métodos:** Dos etapas fueron completadas para lograr el objetivo de investigación. En primer lugar, se desarrolló un set de ítems basados en los "tres componentes centrales del estigma: creencia, emoción y conducta". Se les solicitó a tres jueces independientes evaluar estas preguntas de acuerdo con su contenido y validez. En segundo lugar, siguiendo un diseño instrumental y longitudinal con muestreo no probabilístico por cuotas por género (N = 607) se evaluó la validez y confiabilidad de la escala final. **Resultados:** Una escala unidimensional compuesta por 22 ítems mostró buenos límites estadísticos. Las saturaciones de factores observadas sugieren que los ítems representan adecuadamente la dimensión ($\lambda > .6$), y las estimaciones de confiabilidad son óptimas ($\alpha > .8$; $\omega > .8$). Los resultados sugieren que la escala se puede utilizar independientemente del género de los encuestados.

Palabras clave: stigma; esquizofrenia; validación de escala.



Introduction

Stigma and prejudice are two related and different concepts (Link & Phelan, 2001). Both are complex concepts that involve individual experience, the interaction between marginalised and non-marginalised groups, and social phenomena, including power relations, contingencies, social practices and program policy designs (Allport, 1979; Goffman, 1963). The classic starting point for conceptualise stigma of mental issues is proposed by Goffman (1963, p. 11), as “an attribute that links a person to an undesirable stereotype, leading other people to reduce the bearer from a whole and usual person to a tainted, discounted one”. Hence, stigma is a set of cognitions and feelings, usually negative, against a minority social group with a different condition than the rest of the society (López et al., 2008). On the other hand, in Allport words, prejudice is defined as “an aversive or hostile attitude toward a person who belongs to a group, simply because he belongs to that group, and is therefore presumed to have the objectionable qualities ascribed to the group” (Allport, 1979, p. 7). Therefore, prejudice is a negative attitude against an individual who belongs to a group with a negative social value (Navarro et al., 2012).

Based on these previous works, stigma and prejudice have similarities including, exposure to adverse attitudes, unfair treatment and violence against individuals belonging to disadvantage social groups. Notwithstanding these similarities, stigma and prejudice are different attitudinal concepts (Ottati et al., 2005). Indeed, stigma research has primarily focused on individuals with atypical conditions, including mental illness (e.g., schizophrenia). Conversely, prejudice research focused on the far more common but powerful implications of race, nationality, gender and class division, among other similar variables (Stubber et al., 2008). Therefore, prejudice research has been associated with social processes determined by domination, while stigma research has been related to processes driven by disease avoidance (Phelan et al., 2008).

Three critical components integrate stigma (Reynders et al., 2014). The first one is related to the *cognitive* component of stereotypes. Stereotypes are cognitive structures which helps individual to simplify their context (Corrigan & O’Shaughnessy, 2007). As such, stereotypes towards individuals with mental health issues, including schizophrenia, results in negative beliefs such as *individuals with psychological issues being inept* (Reynders et al., 2014). The second component is prejudice, which represents the agreement of most individuals with these stereotypes or the experience of adverse *affective* reactions, including *anger* or *fear* toward the stigmatised group (Pascal et al., 2023; Reynders et al., 2014). Prejudice

leads to discrimination as a behavioural consequence (Fiske et al., 2010). As such, the third component is the *behavioural* responses to prejudice result in the need for social distance (Rusch et al., 2005). A central concept in the process of stigmatisation is labelling (Reynders et al., 2014). Through labelling, the negative stereotype of being diagnosed with a mental health illness will be triggered (Angermeyer & Matschinger, 2003).

Regarding mental health illness, research has proposed that individuals with typical or common mental health illnesses, such as depression or anxiety, are more accepted and suffer less stigma (Cazzaniga & Suso, 2015). Conversely, individuals diagnosed with schizophrenia are the most who suffer stigma and are less considered by societies (Rusch et al., 2005; Soler, 2019). It is suggested that stigma against individuals diagnosed with schizophrenia is based on mistaken ideas about their dangerousness and violent behaviour (Chang et al., 2018). Research has flagged out that individuals diagnosed with schizophrenia may suffer one or a combined type of stigma: direct discrimination (e.g., refusal to hire the individual) (Grandon et al., 2018; Link et al., 2001), structural discrimination (e.g., less availability of adequate treatment) (Gren et al., 2020 van Zelst, 2008), and social-psychological process which involve the stigmatised individuals' perceptions (Link & Phelan, 2001; Valery & Proteau, 2020).

Stigmatisation may harm the course and the onset of schizophrenia (Graves et al., 2005; van Zelst, 2008). Literature has highlighted several negative consequences impacting individuals diagnosed with schizophrenia. For example, it has been proposed that stigma represents a significant barrier to accomplishing the social integration of these individuals (López et al., 2008; Soler, 2019). Experiences of stigma can be covering almost all aspects of standard of living, including general health and psychiatric treatment (Gonzalez-Torres et al., 2007). Stigma can significantly reduce individuals' self-esteem and life satisfaction (Dickerson et al., 2002; Link et al., 2001; Markowitz, 1998). Stigma produces a vicious cycle of discrimination and disadvantage, which in turn lead individuals into social seclusion (including homelessness), drug or alcohol abuse, reduced job opportunities and excessive institutionalisation. Moreover, stigma does not also negatively impact individuals diagnosed with schizophrenia, evidence support that families or their caregivers suffer social isolation (Caqueo-Urizar et al., 2014). Finally, disadvantages in social legislation and the health care system reflect the harmful effects of stigmatization of individuals diagnosed with schizophrenia (Gaebel & Baumann, 2003). All of these consequences decrease the likelihood of recovery (Gaebel & Baumann, 2003).

In response to the negative impact of the stigma of individuals diagnosed with schizophrenia, research and practitioners focused on developed initiatives to reduce this stigmatization are gaining acceptability. Broad knowledge and successful intervention is the Open the Doors program developed by the World Psychiatric Association, implemented in more than 20 countries to reduce stigma against individuals diagnosed with schizophrenia (Hochman, 2007). For example, international evidence has supported that based interventions research has been successful in Australia (Rosen et al., 2000), Germany (Gaebel & Baumann, 2003), New Zealand (Vaughan & Hansen, 2004), the United Kingdom (Curran, 2004) and Japan (Desapriya & Nobutada, 2002).

Measuring the success of these interventions research relies on quantifying differences between scales over time or between interventions and controls groups over time in a between-within design (McEldoon et al., 2012). Indeed, selecting appropriate and valid scales is critical in designing valid and useful intervention programs and outcomes research (An et al., 2020; Coster, 2013). Therefore, a valid scale is required to ensure the quality of antistigma interventions against individuals diagnosed with schizophrenia in any country or society.

In Chile is estimated that the proportion of individuals diagnosed with schizophrenia is between 1.4 to 4.6 individuals per 1,000 inhabitants (Ministerio de Salud Chile, 2009; World Health Organization, 2017). Furthermore, it is expected that twelve new cases per 100,000 inhabitants will be diagnosed in any given year (Hospital Clínico de la Universidad de Chile, 2016). As a result, schizophrenia treatment is a critical concern and policy priority for Chilean public health (Langer et al., 2017). However, stigma against individuals diagnosed with schizophrenia is a barrier to accessing adequate treatment for these individuals (Grandón et al., 2015). Indeed, in Chile, both the lack of knowledge about schizophrenia and stigma against these individuals, are the main reasons to delay or avoid medical treatment (Vicente et al., 2007).

Notwithstanding local interventions have been developed to reduce stigma against individuals diagnosed with schizophrenia (Solano & Vasquez, 2014), there is a lack of adequate scales to evaluate its efficacy. In Chile, are available three scales to assess stigma against individuals diagnosed with severe mental disorder, such as schizophrenia. The first scale was adapted from *The Community Attitudes toward Mental Illness, CAMI* (Taylor & Dear, 1981) by Grandón et al. (2016). 10 items were integrated into the scale in its final version. Each item is rated using a 5-point Likert scale (1=strongly disagree, 6=strongly agree), that were equally distributed into two factors: acceptance and rejection of mental health centers in the communi-

ty. The authors reported adequate values of Cronbach's alpha $\alpha = 0.69$ for the scale. Although the statistical features reported, this scale has some limitations. First, the scale was designed to assess the behavioural answer of the contestants towards individuals with severe mental disorders. Thus, the scale was not designed to evaluate stigma towards individuals with schizophrenia specifically. Finally, some methodological issues have been reported, including lack of adjustment to the original structure of the scale, bias of the non-probabilistic sample using a cross sectional design and the scale showed differences with respect to the original instrument, particularly in the factorial structure (Grandón et al., 2016).

The second scale was adapted from the *Social Distance Scale* (Bogardus, 1925) by Grandón et al. (2015). Five items were integrated into the scale, each item is rated using a 7-point Likert scale (1 = strongly disagree, 7 = strongly agree). Two factors integrate the scale: closeness and social interaction (2 items) and intimacy and trust (three items). The authors reported acceptable values of Cronbach's alpha $\alpha = 0.78$ for the scale. Notwithstanding this value, the scale has some limitations. First, the developed scale was designed to assess stigma towards individuals with several mental disorders in general, but not the schizophrenia in particular. Hence the scale does not consider the affective and cognitive components of stigma. Finally, the scale has bias of the non-probabilistic sample using a cross sectional design and does not consider the test-retest reliability that allows to account for the temporal stability of the measures (Grandón et al., 2015).

The third scale was adapted from the scale used by the World Health Organisation, WHO in the implementation of the open the doors programme (Stuart et al., 2001), in its Spanish language version adapted in Argentina (Leiderman et al., 2011) by Grandón et al. (2018). This scale was developed to assess stigma towards individuals with schizophrenia by measuring knowledge and level of information with variables which, in other cultural contexts, have been shown to influence the stigma the population attaches to people with schizophrenia.

11 items were integrated into this scale, eight items have a yes/no/don't know answer format; the ninth item investigates the causes of schizophrenia, gives the multiple choice options: genetic, environmental or both; the tenth item asks for knowledge about more effective treatments, gives the options: medication, psychotherapy or both, and the last item investigates if the medical treatment helps them to remain stable, with the choice being yes/no/don't know. Two factors integrate the scale: beliefs on the knowledge of schizophrenia and attitudes towards schizophrenia. The authors reported acceptable values of Cronbach's alpha for each factor $\alpha = 0.72$ and $\alpha = 0.52$. However, the Cronbach's alpha for the whole scale was

not reported. The study has some limitations. First, the study was the non-probability sampling. Furthermore, the authors reported that the scale did not reflect the knowledge well, as a consequence, further research therefore needs to be carried out on the structure of the scale

Furthermore, the three research presented above, reported issues with the translation procedure and the cultural representations of the scales. Regarding to this practice, the literature has flagged two main issues in scales translation. First, the translation from the original language to another represents a methodological problem (Hachey et al., 1995). Second, the cultural representations of the scale might be not a good representation in different settings (Ramada et al., 2013). These related issues are based on the quality of the translation and the comparability of different cultural or ethnic groups (Sperber, 2004).

Consequently, it is not available in Chile appropriate and valid scales to evaluate the stigma against individuals diagnosed with schizophrenia. This study aims to develop a valid scale to evaluate the stigma against individuals diagnosed with schizophrenia within the Chilean population to address this gap.

Method

Stage 1: Generating and Judging Measurement Items

In stage 1, a pool of items was developed based on the three critical components of stigma, *cognitive*, *affective* and *behavioural* answers against individuals diagnosed with schizophrenia. A balance of items between the three components was achieved by considering a potential loss of 75%. In order to further assess the content aspects of content validity, the items were rated by three external expert judges (Netemeyer et al., 2003), which were independent of those who developed the items (Boateng et al., 2018), and they are experts in areas of social psychology and instrument validation. The expert judges were presented with all the items alongside definitions of the three components of stigma (cognitive, affective and behavioural). They were asked to (a) assign each item to whichever component they felt most relevant, b) rate each item from 1 (not adequate) to 3 (clearly adequate) in terms of how well they felt it represented the component that they had selected. Furthermore, the expert judges were asked to provide qualitative feedback on item readability, wording, clarity and overlap. Items were maintained if classified in the original component and rated as clearly representative unanimously by the expert judges. As a result, minor wording changes were implemented to

enhance understanding of several items, and a total of 61 items were retained (28 cognitive, 12 affective and 21 behavioural).

Stage 2: Exploratory Factor Analysis

Sample and Procedure

This research was conducted following an instrumental and longitudinal design. A non-probabilistic with a quota sampling design by gender was chosen. Participation in the study was voluntary and written informed consent was obtained from the study participants. Participants were asked to complete one online survey hosted on Google Forms. A sample compromised by 607 participants was achieved. The sample had approximately twice as many females ($N = 412$, 67.9%) as males ($N = 194$, 32%). Participant ages ranged from 18 to 71 ($M = 32$, $SD = 11,6$). The degree of relationship with individuals diagnosed with schizophrenia was asked of participants. There were 244 (40.2%) participants who reported no direct contact with individuals diagnosed with schizophrenia, 204 (33.6%) participants reported occasional and distant contact with them, 95 (15.7%) participants reported occasional and close contact with these individuals, and 64 (10.5%) participants have had frequent and close relations with individuals diagnosed with schizophrenia. Finally, participants' mental health situation was asked. There were 199 (32.8%) participants who had been diagnosed with any mental health issue during their life, and 408 (67.2%) participants who had been not diagnosed with any mental health issue during their life. A detailed breakdown of sample demographics is presented in Table 1.

Table 1. *Demographic characteristics of the sample*

| | N | % |
|---|-----|-------|
| Marital Status | | |
| Single | 258 | 42,9% |
| Married | 130 | 21,6% |
| Unmarried Couple (Cohabitation Agreement) | 99 | 16,4% |
| Romantical Relationship (No-cohabitation agreement) | 72 | 12% |
| Education Level | | |
| High School | 111 | 18,4% |
| Technical | 111 | 18,4% |
| Undergraduate Degree not completed | 166 | 27,5% |
| Undergraduate Degree | 134 | 22,2% |
| Socioeconomic Status | | |
| Middle | 339 | 56,2% |
| Low | 161 | 26,7% |

Statistical analyses

Firstly, reliability analyses and corrected homogeneity indices were carried out through the Jamovi statistical program. Table 2 is presented the corrected homogeneity index of the set of items of the cognitive dimension. Table 3 is presented the corrected homogeneity index of the set of items of the affective dimension. Finally, Table 4 is presented the corrected homogeneity index of the set of items of the behavioural dimension.

Table 2. *Corrected homogeneity index of cognitive dimension items.*

| Item | Item-rest correlation | Cronbach's α | McDonald's ω |
|-------|-----------------------|---------------------|---------------------|
| COG1 | -0.364 | 0.814 | 0.841 |
| COG2 | 0.449 | 0.773 | 0.817 |
| COG3 | 0.646 | 0.764 | 0.807 |
| COG4 | 0.571 | 0.767 | 0.811 |
| COG5 | 0.475 | 0.772 | 0.814 |
| COG6 | 0.418 | 0.777 | 0.815 |
| COG7 | 0.515 | 0.771 | 0.813 |
| COG8 | 0.384 | 0.778 | 0.817 |
| COG9 | -0.075 | 0.801 | 0.832 |
| COG10 | 0.641 | 0.762 | 0.808 |
| COG11 | 0.008 | 0.796 | 0.830 |
| COG12 | 0.216 | 0.785 | 0.824 |
| COG13 | 0.317 | 0.780 | 0.821 |
| COG14 | 0.263 | 0.783 | 0.822 |
| COG15 | -0.212 | 0.805 | 0.838 |
| COG16 | 0.608 | 0.767 | 0.808 |
| COG17 | -0.072 | 0.798 | 0.834 |
| COG18 | 0.561 | 0.769 | 0.810 |
| COG19 | 0.548 | 0.771 | 0.810 |
| COG20 | 0.324 | 0.779 | 0.822 |
| COG21 | 0.310 | 0.780 | 0.823 |
| COG22 | 0.612 | 0.766 | 0.809 |
| COG23 | 0.566 | 0.766 | 0.810 |
| COG24 | 0.356 | 0.779 | 0.818 |
| COG25 | 0.391 | 0.776 | 0.819 |
| COG26 | 0.288 | 0.781 | 0.823 |
| COG39 | 0.098 | 0.791 | 0.828 |
| COG40 | 0.230 | 0.784 | 0.823 |

Table 3. *Corrected homogeneity index of affective dimension items.*

| Item | Item-rest correlation | Cronbach's α | McDonald's ω |
|-------|-----------------------|---------------------|---------------------|
| EMO27 | 0.666 | 0.768 | 0.798 |
| EMO28 | 0.608 | 0.776 | 0.803 |
| EMO29 | 0.706 | 0.763 | 0.793 |
| EMO30 | 0.225 | 0.812 | 0.834 |
| EMO31 | 0.693 | 0.765 | 0.795 |
| EMO32 | 0.411 | 0.792 | 0.820 |
| EMO33 | 0.393 | 0.794 | 0.821 |
| EMO34 | 0.338 | 0.800 | 0.825 |
| EMO35 | 0.018 | 0.832 | 0.845 |
| EMO36 | 0.567 | 0.781 | 0.806 |
| EMO37 | 0.638 | 0.769 | 0.802 |
| EMO38 | 0.305 | 0.802 | 0.828 |

Table 4. *Corrected homogeneity index of behavioural dimension items.*

| Item | Item-rest correlation | Cronbach's α | McDonald's ω |
|-------|-----------------------|---------------------|---------------------|
| CON41 | 0.577 | 0.901 | 0.903 |
| CON42 | 0.512 | 0.902 | 0.905 |
| CON43 | 0.643 | 0.899 | 0.902 |
| CON44 | 0.687 | 0.898 | 0.901 |
| CON45 | 0.660 | 0.898 | 0.901 |
| CON46 | 0.620 | 0.899 | 0.902 |
| CON47 | 0.677 | 0.898 | 0.901 |
| CON48 | 0.558 | 0.901 | 0.904 |
| CON49 | 0.628 | 0.899 | 0.902 |
| CON50 | 0.499 | 0.902 | 0.905 |
| CON51 | 0.551 | 0.901 | 0.904 |
| CON52 | 0.090 | 0.912 | 0.913 |
| CON53 | 0.645 | 0.899 | 0.902 |
| CON54 | 0.555 | 0.901 | 0.904 |
| CON55 | 0.480 | 0.903 | 0.906 |
| CON56 | 0.245 | 0.907 | 0.910 |
| CON57 | 0.412 | 0.904 | 0.907 |
| CON58 | 0.571 | 0.901 | 0.904 |
| CON59 | 0.539 | 0.901 | 0.904 |
| CON60 | 0.484 | 0.903 | 0.906 |
| CON61 | 0.543 | 0.901 | 0.904 |

It was decided to discard those items that presented inverse correlations, or less than 0.3, with the corrected total score. As a result, ten items were eliminated iteratively. Secondly, to assess the internal validity of the scale, four models were contrasted following the sequential analyses:

- a) An exploratory structural equation model (ESEM) with GEOMIN rotation based on the polychoric correlation matrix was selected. Furthermore, the robust weighted least squares estimation method (WLSMV) (Asparouhov & Muthén, 2009) was considered. The initial model of three dimensions is presented in Figure 1. The general fits of the model were evaluated according to

the criteria proposed by Schreiber (2017) (e.g. CFI > .95, TLI > .95 and RMSEA < .06).

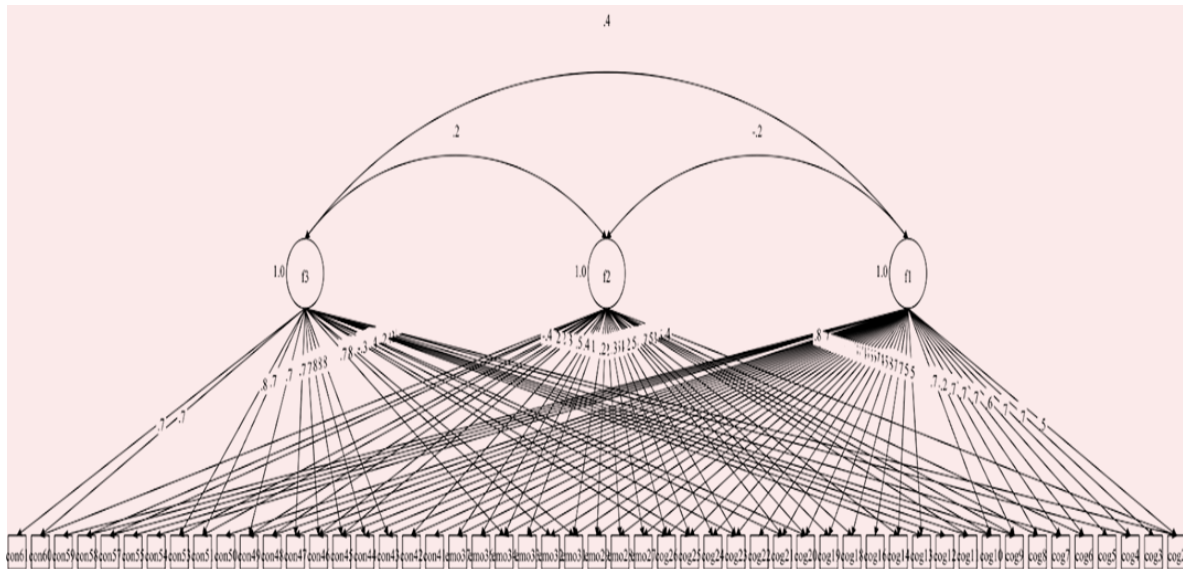


Figure 1. Initial ESEM model composed of three dimensions F1 = cognitive; F2 = affective; F3 = behavioural

b) Subsequently, given the presence of high cross-loads, a bifactor model was tested with three specific factors (cognitive, affective and behavioural) and a general factor (Reise et al., 2013). In Figure 2, this model is presented.

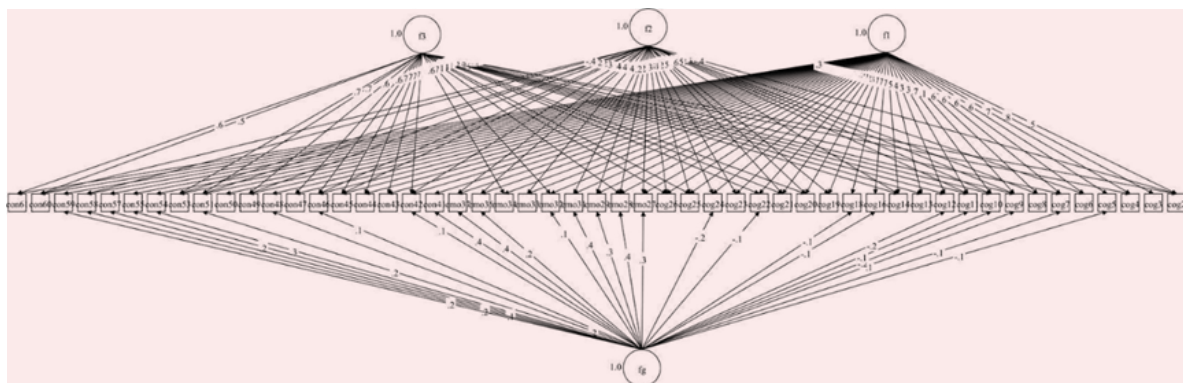


Figure 2. ESEM bifactor model, Fg = general factor; F1 = cognitive; F2 = affective; F3 = behavioural

c) However, the bifactor model did not present coherent loadings with the proposed structure, with a predominance of the general factor, over the specific factors. Thus, it was decided to test a one-dimensional solution by confirmatory factor analysis (CFA), which initially contains 51 items (Figure 3). Furthermore, the WLSMV based on the polychoric correlation matrix was considered. This method was iteratively debugged based on two criteria: re-

tention of strong factorial loads ($\lambda > .5$) and elimination of redundant elements. As a results, the final model (Figure 4) was integrated by 22 items.

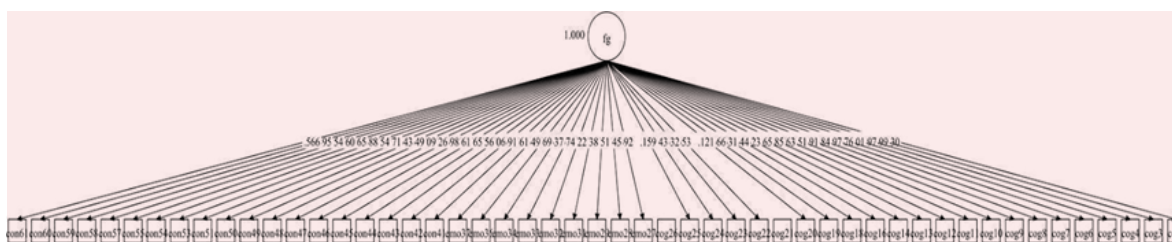


Figure 3. One-dimensional confirmatory factor analysis (CFA) with 51 items

- d) An invariance test was carried out in the one-dimensional model of 22 items to evaluate the stability of the scale. It was considered constant changes in CFI less than 0.010 and RMSEA less than 0.015 (Chen, 2007). This final model is presented in Figure 4. After that, reliability analyses were carried out for the final scale. These analyses were conducted with Mplus 8.2 (Muthén & Muthén, 2017). It is noticed that the factorial loadings of the items are significant to the general factor.

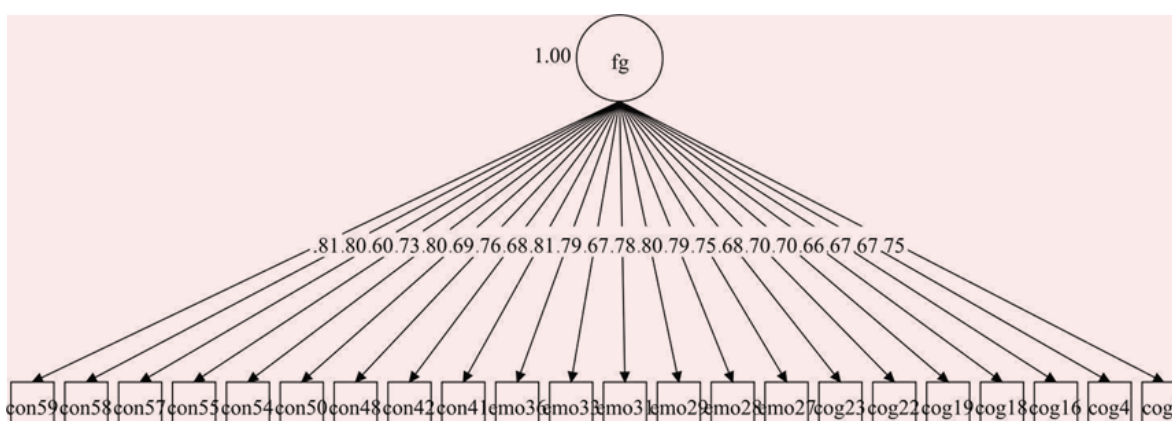


Figure 4. Final model. One-dimensional confirmatory factor analysis (CFA) of 22 items

Results

In Table 5, the statistical indicators of the goodness of fit analyses for the scale used by considering four models are presented. First, values of the initial model, Model 1, are partially adjusted to what is recommended ($CFI > .95$; $TLI > .95$; $RMSEA < .06$) (Schreiber, 2017). Notwithstanding that all the contrasted models fit as an explanation of the covariance variance matrix, the ESEM model presented interpretation problems, including items redundancy, reducing factor loadings, and cross loads presence. Furthermore, a great percentage of cross-loadings were observed. Second, Model 2 is a bi-factorial model with an adequate fit. However, it was not

possible to establish a clearly defined structure in its factorial loadings while the general factor loads indicated the one-dimensional trend of the scale. Third, Model 3 presented important issues in its fit standards. Thus, it was necessary to debug it iteratively. Items compatible with the one-dimensional model were kept. Finally, Model 4 shown an adequated explanation of observed relations, with all the items with relevant factorial loadings and absence of relevant covariations between item errors.

Table 5. *Global fit indicators of measurement models*

| Model | Pairs Number | x2 | DF | P | TLI | CFI | RMSEA | RMSEA CI 90% | |
|-----------------------|--------------|-----------|------|------|------|------|-------|--------------|------|
| | | | | | | | | Low | Upp |
| Model 1 (51 items) | 303 | 2030.784 | 1125 | .000 | .973 | .976 | .036 | .034 | .039 |
| Model 2 (51 items) | 351 | 1603.661 | 1077 | .000 | .983 | .986 | .028 | .025 | .031 |
| Model 3 (51 items) | 204 | 13315.211 | 1224 | .000 | .665 | .678 | .128 | .126 | .130 |
| Model 4 (22 items) | 88 | 937.449 | 209 | .000 | .955 | .959 | .077 | .072 | .082 |

Notes: Model 1= Initial ESEM 51 items; Model 2= ESEM bi-factorial 51 items; Model 3= unidimensional CFA of 51 items; Model 4= Unidimensional CFA of 22 items; x2= Chi-square; DF= Degrees of freedom; P= Statistical significance; CFI = Comparative fit index; TLI = Tucker-Lewis index; RMSEA = Root mean square error of approximation; CI= Confidence interval.

In table 6, the standardised factorial loadings of the final model (One-Dimensional, 22 Items) and the reliability coefficients' estimates. The observed factor loadings suggest that the items adequately represent the dimension ($\lambda > .6$). Moreover, the reliability estimates are optimal ($\alpha > .8$; $\omega > .8$).

| <i>Escala de estigmatización hacia personas diagnosticadas con esquizofrenia</i> | |
|---|------|
| Suelen ser un peligro para la sociedad | .751 |
| Deberían estar hospitalizadas | .666 |
| Suelen ser perversos | .673 |
| No deberían trabajar | .658 |
| Suelen ser ineficientes | .703 |
| Suelen ser potenciales asesinos | .698 |
| No deberían tener hijos | .682 |
| Me hacen sentir temor | .752 |
| Me hacen sentir rechazo | .791 |
| Me hacen sentir inseguro si están cerca de mí | .802 |
| Me incomodan | .779 |
| Me hacen sentir vergüenza | .669 |
| Me hacen sentir desagrado | .794 |
| Son personas que trato de evitar | .813 |
| Rechazaría un familiar con esquizofrenia | .679 |
| Los ignoro todo lo que puedo | .764 |
| No daría trabajo a alguien con esquizofrenia | .690 |
| Evité hablarles | .800 |
| Cruzaría la calle si se me acercaran | .732 |
| Los golpearía si me sintiera en peligro | .604 |
| Son personas que evitaría tocar | .800 |
| Intentó evitar que se acerque a mis hijos | .805 |
| Alpha de Cronbach | .941 |
| Coeficiente Omega | .942 |

Finally, in Table 7, the invariance test results according to the sex of the participants are presented.

Table 7. *The contrast of invariance measure to the final scale (22 items) between men and women*

| | Pairs Number | χ^2 | DF | P | CFI | RMSEA | $\Delta\chi^2$ | ΔDF | $P_{\Delta\chi^2}$ | ΔCFI | $\Delta RMSEA$ |
|------------|--------------|----------|-----|------|------|-------|----------------|-------------|--------------------|--------------|----------------|
| Configural | 176 | 1121.821 | 418 | .000 | .965 | .076 | | | | | |
| Metric | 155 | 1148.337 | 439 | .000 | .965 | .074 | 35.372 | 21 | .025 | 0 | -.002 |
| Scalar | 112 | 1109.269 | 482 | .000 | .969 | .066 | 116.041 | 64 | .000 | .004 | -.01 |

Notes: Pairs number= parameters number; χ^2 = Chi-square; DF= Degrees of freedom; P= Statistical significance; $\Delta\chi^2$ = Chi-square difference; Δgl : Degrees of freedom difference; Δp : Statistical significance difference.

The statistical differentials between the model without restrictions (e.g. configuration) and models with metric and scalar restrictions allow the invariance.

Discussion

This study aimed to develop a valid scale of stigma against people with a schizophrenia diagnosis to fill a gap in Chilena research about this topic. As mentioned, previous valuable research (Grandón et al., 2015; Grandón et al., 2016;

Grandón et al., 2018) has attempted to fill this gap by adapting international scales. However, methodological issues, like fit indexes under the suggested standards in the literature (Scheiber, 2017), have been reported in these studies. Furthermore, adapting scales practices might have other problems, including translation quality and cultural representations (Hachey et al., 1995; Ramada et al., 2013; Sperber, 2004). Therefore, in this research, these previous limitations have been considered to build and validate an original scale of stigma against people with a schizophrenia diagnosis in the Spanish language and by considering the local cultural representation.

At the beginning of this research, a tridimensional model was proposed to assess stigma against people with a schizophrenia diagnosis. Cognitive, affective and behavioural components are integrated into this model. However, the interrelationships between these components, and the ability of individuals to make differentiated judgments, led us to propose a one-dimensional model that jointly incorporates the three attitudinal concepts. Previous research has supported this decision, which has explained that cognitive, affective and behavioural components are strongly interrelated and mutually reinforcing (López et al., 2008; Reynders et al., 2014). It means that stereotypes (cognitive) are accepted if they are concordant with emotional prejudice (affective), resulting in discriminatory acts (behaviour).

Finally, a one-dimensional model was considered to develop a scale to assess stigma against people with a schizophrenia diagnosis. Twenty-two items integrate the final version of this one-dimensional scale. The statistical results showed that this scale has adequate psychometric properties, including strong reliability and validity evidence based on the internal structure. Furthermore, the invariance tests allow accepting the equivalence between men and women, which means that this scale can be used the respondents' gender irrespectively.

This research has limitations to be mentioned. First, this study was developed in the context of the COVID-19 pandemic. This context might impact the rates of answers and on the psychological-emotional condition of the respondents. Second, the study participants in their majority are from one Chilean city (Arica), which might imply a non-representative of the total Chilean population. Future research may apply this scale within different populations and settings to have more comprehensive results and further evidence about its reliability and internal consistency.

Conclusion

The relevance of this research is to build and validate an original scale of stigma against people with a schizophrenia diagnosis in the Spanish language and by considering the local cultural representation. It is a short one-dimensional short scale (22 items) that were considered the three critical components of the stigma: cognitive, affective and behavioural. This scale can be used for research purposes and to understand better the stigma against people with a schizophrenia diagnosis in Chile. Furthermore, applying this scale in different settings and participants can provide further insights into its psychometric properties.

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